

Nonparametric Estimation of Expected Shortfall

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ABSTRACT The expected shortfall of a random variable is a popular risk measure. In this paper, we propose a new nonparametric estimator of the expected shortfall of a random variable. The estimator is based on the empirical distribution function of the random variable. We prove that the estimator is consistent and asymptotically normal. Monte Carlo simulations show that the estimator performs well in finite samples. The proposed estimator is applied to the expected shortfall of a random variable. The results show that the estimator is accurate and efficient.

1. INTRODUCTION

Let $\{X_t\}_{t=1}^n$ be a sequence of independent random variables with common distribution P on $\mathbb{R}$. Let $\{Y_t\}_{t=1}^n$ be a sequence of random variables such that $Y_t = -\log X_t/X_{t-1}$ and $\{Y_t\}_{t=1}^n$ is independent of $\{X_t\}_{t=1}^n$. Let F be the cumulative distribution function of Y_1 . Then for any $p \in (0, 1)$ we have

$$\nu_p = \inf\{y \in \mathbb{R} : F(y) \geq 1 - p\} \tag{1}$$

where ν_p is the p -quantile of Y_1 . Let F be the cumulative distribution function of

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in the following, we denote the nonfunctional dependence of the observed data on the functional covariate by μ_p and the conditional density of Y_t given I_t by f_t .

Let p be a fixed number in $(0, 1)$ and let ν_p be a positive number such that $\nu_p \rightarrow 0$ as $n \rightarrow \infty$ and $n\nu_p \rightarrow \infty$. Let $Y_{(r)}$ be the r -th order statistic of $\{Y_t\}_{t=1}^n$. A functional density f is said to be p -smooth if

2. NONPARAMETRIC ESTIMATORS

A nonparametric estimator of the functional density f of Y_t is defined as $\hat{f}_n$ such that $\hat{f}_n(Y_{(r)}) = Y_{([n(1-p)]+1)}$ for $r \leq p$ and $\hat{f}_n(Y_{(r)}) = 0$ for $r > p$. The estimator $\hat{f}_n$ is said to be p -smooth if

$$\mu_p \left[\frac{\sum_{t=1}^n Y_t I_t \geq \nu_p}{\sum_{t=1}^n I_t Y_t \geq \nu_p} \right] \leq np \quad \text{and} \quad \sum_{t=1}^n Y_t I_t \geq \nu_p \quad (4)$$

Let I_t be a functional covariate.

Let ν_p be a positive number such that $\nu_p \rightarrow 0$ as $n \rightarrow \infty$ and $n\nu_p \rightarrow \infty$. Let K be a kernel function satisfying $\int_{-\infty}^{\infty} K(u) du = 1$ and $G_h(t) = \int_{-\infty}^{\infty} K(u) f(t+u) du$ and $G_h(t) = G(t)/h$ for $h > 0$. Let $\nu_{p,h}$ be a positive number such that $\nu_{p,h} \rightarrow 0$ as $n \rightarrow \infty$ and $n\nu_{p,h} \rightarrow \infty$. Let $S_h(z)$ be a functional density estimator of f such that

$$S_h(z) = n^{-1} \sum_{t=1}^n G_h(z - Y_t) \quad (4)$$

A nonparametric estimator of ν_p is denoted by $\hat{\nu}_{p,h}$ and is defined as $\hat{\nu}_{p,h} = p$ if $S_h(z) \geq \nu_p$ and $\hat{\nu}_{p,h} = 0$ otherwise. By the functional density f and ν_p we define G_h and $\nu_{p,h}$ respectively. Let $\nu_{p,h}$ be a positive number such that $\nu_{p,h} \rightarrow 0$ as $n \rightarrow \infty$ and $n\nu_{p,h} \rightarrow \infty$. Let $S_h(z)$ be a functional density estimator of f such that

$$\mu_{p,h} \left[np \sum_{t=1}^n Y_t G_h(\nu_{p,h} - Y_t) \right] \leq np \quad (4)$$

²Its statistical properties and how to obtain the standard errors are considered in Chen and Tang (2005). See also Cai (2002) and Fan and Gu (2003) for kernel estimators.

By definition, $\nu_{p,h}$ is a probability measure on $\mathcal{F}_p$. For $h \geq 1$, let ν_p be the probability measure on $\mathcal{F}_p$ defined by $\nu_p(A) = \sum_{h=1}^{\infty} \nu_{p,h}(A)$. For $p \geq 1$, let ν_p be the probability measure on $\mathcal{F}_p$ defined by $\nu_p(A) = \sum_{h=1}^{\infty} \nu_{p,h}(A)$.

3. MAIN RESULTS

Let ν_p be a probability measure on $\mathcal{F}_p$. For $p \geq 1$, let ν_p be the probability measure on $\mathcal{F}_p$ defined by $\nu_p(A) = \sum_{h=1}^{\infty} \nu_{p,h}(A)$.

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$$\alpha_k = \inf_{A \in \mathcal{F}_1, B \in \mathcal{F}_{k+1}} |P(AB) - P(A)P(B)|.$$

Let ν_p be a probability measure on $\mathcal{F}_p$. For $p \geq 1$, let ν_p be the probability measure on $\mathcal{F}_p$ defined by $\nu_p(A) = \sum_{h=1}^{\infty} \nu_{p,h}(A)$.

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$$K = \int_{-1}^1 u^2 K(u) du, \quad \sigma_K^2 > 0, \quad \text{and } H = \int_{-1}^1 u^2 K(u) du, \quad \text{and } L = \int_{-1}^1 u^2 K(u) du.$$

Condition 4.1 is satisfied. For $p \geq 1$, let ν_p be the probability measure on $\mathcal{F}_p$ defined by $\nu_p(A) = \sum_{h=1}^{\infty} \nu_{p,h}(A)$.

[illegible]

Let μ_p denote the population mean of μ_p . Under condition 4, as $n \rightarrow \infty$

$$\sqrt{pn}\sigma_0^{-1}(\mu_{p,h} - \mu_p) \xrightarrow{d} N(0, 1) \quad (4)$$

and furthermore

$$Bias(\mu_{p,h}) = -\frac{1}{2}p^{-1}\sigma_K^2 f(\nu_p) + o(p^{-2}) \quad (4)$$

$$Var(\mu_{p,h}) = p^{-1}n^{-1}\sigma_0^2 + o(n^{-1}) \quad (4)$$

By comparing the above two conditions, we see that the proposed $\mu_{p,h}$ is more efficient than the usual μ_p in the sense that the variance of $\mu_{p,h}$ is smaller than that of μ_p . In fact, the variance of $\mu_{p,h}$ is of order n^{-1} , while the variance of μ_p is of order $n^{-1}p$. The bias of $\mu_{p,h}$ is of order p^{-1} , while the bias of μ_p is of order $p^{-1}n$. Therefore, $\mu_{p,h}$ is a more efficient estimator than μ_p . The density of $\mu_{p,h}$ is denoted by $f_{\mu_{p,h}}$.

The second part of the proof is concerned with the asymptotic normality of $\mu_{p,h}$. It is well known that the asymptotic normality of $\mu_{p,h}$ is equivalent to the asymptotic normality of μ_p . In fact, the asymptotic normality of μ_p is equivalent to the asymptotic normality of $\mu_{p,h}$. Therefore, the asymptotic normality of $\mu_{p,h}$ is equivalent to the asymptotic normality of μ_p . The asymptotic normality of μ_p is well known. Therefore, the asymptotic normality of $\mu_{p,h}$ is well known.

It is also noted that the above two conditions are satisfied by the proposed $\mu_{p,h}$. Therefore, the proposed $\mu_{p,h}$ is a more efficient estimator than μ_p . The asymptotic normality of $\mu_{p,h}$ is well known. Therefore, the asymptotic normality of $\mu_{p,h}$ is well known.

4. SIMULATION STUDY

In this section, we report the results of a simulation study. The purpose of the simulation study is to compare the performance of the proposed $\mu_{p,h}$ with the usual μ_p . The simulation study is conducted under the following conditions:

Let $\mathcal{A}$ be a σ -algebra on $\mathcal{Y}$ and let $\mathcal{B}$ be a σ -algebra on $\mathcal{X}$.

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$p \leq n/q$ and $\nu^2 q \leq C p$. Applying the Cauchy-Schwarz inequality and the Boole inequality, we have

$$\begin{aligned} & P\{|F_n - \nu_p| \geq C_1 \epsilon_n\} \\ & \leq P\left(-\frac{C_1^2 \epsilon_n^2 q}{\sigma^2}\right) \leq P\left\{-\frac{C_1^2 \epsilon_n^2 q}{\sigma^2} \geq -C_2 \epsilon_n q\right\} \end{aligned} \quad \text{A.4}$$

$$\begin{aligned} & \leq P\left(-\frac{C_1^2 \epsilon_n^2 q}{\sigma^2}\right) \leq P\{-C_2 \epsilon_n q\} \end{aligned} \quad \text{A.4}$$

Since $C_2 > 0$ and $n \epsilon_n^2 \rightarrow \infty$ as $n \rightarrow \infty$, we have $n \epsilon_n q \rightarrow \infty$. Hence, by the lemma, we have

$$\left\{-\frac{C_1^2 \epsilon_n^2 q}{\sigma^2} \geq -C_2 \epsilon_n q\right\} \leq C \epsilon_n^{-1/2} q \rho^{[n^{1/2} \log^{-1}(n)/2]} \quad \text{A.4}$$

Since $n \epsilon_n^2 \rightarrow \infty$ as $n \rightarrow \infty$, we have $n \epsilon_n q \rightarrow \infty$. Hence, by the lemma, we have

Lemma 1. Under the conditions of A.4 and for any $\kappa > 0$,

$$n^{-1} \sum_{t=1}^n |Y_t - \nu_p| I\{Y_t \geq \nu_p\} I\{Y_t \geq \nu_p\} = o_p(n^{-3/4+\kappa})$$

Proof. Let $W_t = Y_t - \nu_p I\{Y_t \geq \nu_p\} - I\{Y_t \geq \nu_p\}$. Then, we have $E W_t = 0$ and

$$\begin{aligned} I_{t1} &= E\{Y_t - \nu_p I\{Y_t \geq \nu_p\} - I\{Y_t \geq \nu_p\} | \nu_p \leq Y_t < \nu_p\} \quad \text{and} \\ I_{t2} &= E\{Y_t - \nu_p I\{Y_t \geq \nu_p\} - I\{Y_t \geq \nu_p\} | \nu_p < Y_t < \nu_p\}. \end{aligned}$$

Let $a \in (0, 1/2)$ and $\eta > 0$.

$$\begin{aligned} I_{t11} &= E\{Y_t - \nu_p I\{Y_t \geq \nu_p\} - I\{Y_t \geq \nu_p\} | \nu_p \geq \nu_p - n^{-a} \eta\}, \\ I_{t12} &= E\{Y_t - \nu_p I\{Y_t \geq \nu_p\} - I\{Y_t \geq \nu_p\} | \nu_p < \nu_p - n^{-a} \eta\}, \\ I_{t21} &= E\{Y_t - \nu_p I\{Y_t \geq \nu_p\} - I\{Y_t \geq \nu_p\} | \nu_p \leq \nu_p + n^{-a} \eta\} \quad \text{and} \\ I_{t22} &= E\{Y_t - \nu_p I\{Y_t \geq \nu_p\} - I\{Y_t \geq \nu_p\} | \nu_p > \nu_p + n^{-a} \eta\}. \end{aligned}$$

Applying the Cauchy-Schwarz inequality, we have

$$|I_{tk1}| \leq \sqrt{E(Y_t - \nu_p)^2 P(|\nu_p - \nu_p| \geq n^{-a} \eta)}$$

em Lemma 2.1 e f.c. $E|\nu_p - \nu_p|^2 = O(n^{-1})$ e p y

$$I_{tk1} \rightarrow 0 \text{ e ponem y f } \quad \text{A 4}$$

o e I_{t12} e no $|I_{t12}| \leq E\{Y_t - \nu_p I_{\nu_p \leq Y_t < \nu_p} n^{-a} \eta\}$. e n

$$I_{t12} \leq \int_{\nu_p}^{\nu_p + n^{-a} \eta} d\mu(z - \nu_p) f(z) dz = O(n^{-2a})$$

em Lemma 2.1 e p p o e c n o $I_{t22} = O(n^{-2a})$ e e nd A 4
e n y o o a $-\gamma$ e e $\gamma > 0$ y

$$E W_t = O(n^{-1+\kappa}) \quad \text{A 4}$$

fo n y p o e κ n n p e

$$E\left[n^{-1} \sum Y_t - \nu_p \{I_{Y_t \geq \nu_p} - I_{Y_t \geq \nu_p}\}\right] = O(n^{-1+\kappa}) \quad \text{A 4}$$

e no con de $Var W_i$ o $a \in \mathbb{R}$, $/ 4$

$$\begin{aligned} E W_t^2 &= E\left[Y_t - \nu_p \{I_{Y_t \geq \nu_p} - I_{Y_t \geq \nu_p} I_{Y_t \geq \nu_p} - I_{Y_t \geq \nu_p}\}\right] \\ &= E\left[Y_t - \nu_p \{I_{\nu_p > Y_t \geq \nu_p} - I_{\nu_p > Y_t \geq \nu_p}\}\right] \\ &= E\left[Y_t - \nu_p I_{\nu_p \leq Y_t < \nu_p} \{I_{\nu_p \geq \nu_p - n^{-a} \eta} - I_{\nu_p < \nu_p - n^{-a} \eta}\}\right] \\ &= E\left[Y_t - \nu_p I_{\nu_p > Y_t \geq \nu_p} \{I_{\nu_p \geq \nu_p - n^{-a} \eta} - I_{\nu_p < \nu_p - n^{-a} \eta}\}\right]. \end{aligned}$$

No e

$$\begin{aligned} E\{I_{\nu_p \leq Y_t < \nu_p} I_{\nu_p \leq \nu_p - n^{-a} \eta}\} &\leq P(|\nu_p - \nu_p| \geq n^{-a} \eta) \text{ nd} \\ E\{I_{\nu_p > Y_t \geq \nu_p} I_{\nu_p > \nu_p - n^{-a} \eta}\} &\leq P(|\nu_p - \nu_p| \geq n^{-a} \eta) \end{aligned}$$

con e e o e ponem y f p ed y Lemma 2.1 App y n C y.
z neq y e

$$\begin{aligned} E\{Y_t - \nu_p I_{\nu_p \leq Y_t < \nu_p} I_{\nu_p \leq \nu_p - n^{-a} \eta}\} &\text{ nd} \\ E\{Y_t - \nu_p I_{\nu_p > Y_t \geq \nu_p} I_{\nu_p \geq \nu_p - n^{-a} \eta}\} & \end{aligned}$$

con $\tilde{e}$ e q ze o e ponem γ y f γ e γ en pp y n γ e γ e γ od γ e γ A γ γ e

$$E\{Y_t - \nu_p \mathbb{I}_{\nu_p \leq Y_t < \nu_p} \mid \nu_p \geq \nu_p - n^{-a} \eta\} = O(n^{-3a}) \quad \text{nd}$$

$$E\{Y_t - \nu_p \mathbb{I}_{\nu_p > Y_t \geq \nu_p} \mid \nu_p < \nu_p - n^{-a} \eta\} = O(n^{-3a})$$

n γ y γ γ $E W_t^2$ γ $n^{-3/2+\kappa}$ γ nd A γ e n $Var W_t$ γ $n^{-3/2+\kappa}$ By γ y od fy n γ γ e de γ on fo $Var W_t$ γ y e γ o n γ fo ny t_1, t_2 $Cov W_{t_1}, W_{t_2}$ γ $n^{-3/2+\kappa}$ γ e efo e

$$Var\left[n^{-1} \sum_{i=1}^n Y_t - \nu_p \mathbb{I}_{Y_t \geq \nu_p} - \mathbb{I}_{Y_t \geq \nu_p}\right] = o(n^{-3/2+\kappa}) \quad \text{A } \gamma$$

γ γ A γ e d y e γ e γ e

LEMMA Le $\beta = np^{-1} \sum Y_t G_h(\nu_p - Y_t)$ nd $\eta = n^{-1} \sum_{i=1}^n Y_t K_h(\nu_p - Y_t)$ nde γ cond γ γ

$$\gamma Cov\left[\beta, \{p - S_h(\nu_p)\} \{f(\nu_p) - f(\nu_p)\}\right] = o(n^{-1})$$

$$\gamma Cov\left[\beta, \eta - \eta \mathbb{I}_{\{p - S_h(\nu_p)\}}\right] = o(n^{-1})$$

$$c \gamma Cov\left[\{p - S_h(\nu_p)\}, \eta - \eta \mathbb{I}_{\{p - S_h(\nu_p)\}}\right] = o(n^{-1})$$

OO e on y p e n γ p oof of γ γ p oof fo γ e γ e De ne $\beta = E \beta$ Le $\beta - \beta = n^{-1} \sum \psi_1 Y_t \mathbb{I}_{\{f(\nu_p) - f(\nu_p)\}} - n^{-1} \sum \psi_2 Y_t = O(n^{-2})$ nd $p - F_h(\nu_p) = n^{-1} \sum \psi_3 Y_t = O(n^{-2})$ fo γ e f nc on ψ_j j , nd $\gamma E\{\psi_j Y_t\} = 0$ n γ nce $\psi_2 Y_t = K_h(\nu_p - Y_t) - E\{K_h(\nu_p - Y_t)\}$ nd $\psi_3 Y_t = G_h(\nu_p - Y_t) - E\{G_h(\nu_p - Y_t)\}$ γ γ pp o γ n B n γ p γ

$$A = |E\left[\beta - \beta \mathbb{I}_{\{p - S_h(\nu_p)\}} \{f(\nu_p) - f(\nu_p)\}\right]|$$

$$\leq n^{-2} \sum_{i \geq 1, j \geq 1, i+j \leq n} |E\{\psi_1 Y_{i+1} \psi_2 Y_i \psi_3 Y_{i+j}\}| = O(n^{-1}) + n^{-2} = O(n^{-1}) \quad \text{A } \gamma$$

γ e e nd c γ γ d e en pe γ on γ e nd ce γ Le $p = \delta$ $q = \delta$ nd $s^{-1} = p^{-1} - q^{-1}$ fo γ e po γ γ δ γ D γ do γ neq γ

$$|E\{\psi_1 Y_{i+1} \psi_2 Y_i \psi_3 Y_{i+j}\}| \leq \|\psi_1 Z_{i+1}\|_p \|\psi_2 Y_i \psi_3 Z_{i+j}\|_q \alpha^{1/s}$$

e e . q e p o y n e o

By Hööörn $r \leq n^a$ for $a \in \mathbb{R}$, and $k' \leq n^c$ for $c \in \mathbb{R}$, $-a < c$ we can choose ϵ and δ of A such that $\epsilon \leq \delta$ for $n \rightarrow \infty$ hence

$$n \xrightarrow{p} \infty \text{ as } n \rightarrow \infty. \quad \text{A} \quad \square$$

We also have $S_{n,1} \leq n^{-1/2} \sum_{j=1}^r W_{j,n} = o_p(1)$ $\square$

By applying the inequality used in Yao et al. (2003) and the concentration of $W_{j,n}$ $\square$

Le. $\eta = E\{p^{-1}Y_tK_h(Y_t - \nu_p)\} = p^{-1} \int (\nu_p - y)K(y)f(\nu_p - y)dy = p^{-1}\nu_p f(\nu_p) = O(h^2)$
 et on a de plus, on a pour $\alpha = 1$ en ce qui concerne la norme L_2 en n Bo. $\eta = E\{np^{-1} \sum Y_tK_h(Y_t - \nu_p) - Y_{t,h} - \nu_p\} = O(n^{-1})$ en ce qui concerne A_1 .

$$E\{np^{-1} \sum Y_tK_h(Y_t - \nu_p) - Y_{t,h} - \nu_p\} = \eta E(Y_{t,h} - \nu_p) = O(n^{-1}) \\ = -\frac{1}{2}p^{-1}\nu_p f'(\nu_p)h^2\sigma_K^2 = o(h^2) = O(n^{-1}) \text{ A } 4$$

Cela ne nous permet pas de conclure.

$$E(\mu_{p,h}) = \mu_p + \frac{1}{2}p^{-1}\sigma_K^2 f''(\nu_p)h^2 = o(h^2) = O(n^{-1})$$

Le lemme est donc établi.

et on a de plus, la variance de $\mu_{p,h}$. Le $A_1 = np^{-1} \sum_{t=1}^n \{Y_tG_h(\nu_p - Y_t) - Y_{t,h} - Y_tK_h(\nu_p - Y_{t,h} - \nu_p)\}$ est une somme de variables indépendantes. On a donc
 en

$$\text{Var}(A_1) = \text{Var}\{np^{-1} \sum Y_tG_h(\nu_p - Y_t)\} = \text{Var}\{\eta(Y_{p,h} - \nu_p)\} \\ = \text{Cov}\{np^{-1} \sum Y_tG_h(\nu_p - Y_t), \eta(Y_{p,h} - \nu_p)\}. \text{ A } 4$$

On a également

$$\text{Var}\{np^{-1} \sum Y_tG_h(\nu_p - Y_t)\} \\ = n^{-1}p^{-2} \left[\text{Var}\{Y_tG_h(\nu_p - Y_t)\} + \sum_{k=1}^{n-1} -k/n \text{Cov}\{Y_1G_h(\nu_p - Y_1), Y_{k+1}G_h(\nu_p - Y_{k+1})\} \right].$$

Le $c_K = \int_{-\infty}^{\infty} yK(y)dy = \int_{-\infty}^u K(y)dy$ et on a

$$\text{Var}\{Y_tG_h(\nu_p - Y_t)\} = \int z^2G_h^2(\nu_p - z)f(z)dz - p^2\mu_p^2 = O(h^2) \\ = \int_{-\infty}^{\infty} K(y)dy \left[\int_{-\infty}^u K(y)dy \left\{ \int_{\nu_p}^{\infty} z^2f(z)dz - \int_{\nu_p-hv}^{\nu_p} z^2f(z)dz \right\} \right. \\ \left. \int_u^{\infty} K(y)dy \left\{ \int_{\nu_p}^{\infty} z^2f(z)dz - \int_{\nu_p-hu}^{\nu_p} z^2f(z)dz \right\} \right] - p^2\mu_p^2 = O(h^2) \\ \text{Var}\{Y_tI(Y_t \geq \nu_p)\} = \nu_p^2f(\nu_p)c_K = O(h^2) \text{ A } 4$$

On a donc et le lemme est établi.

$$\text{Var}\{np^{-1} \sum Y_tG_h(\nu_p - Y_t)\} = p^{-2}\text{Var}\{\phi_1(\nu_p)\} = n^{-1} \nu_p^2f(\nu_p)c_K = o(n^{-1}) \text{ A } 4$$

The second equation on the left hand side of A.4 is

$$\begin{aligned} & \text{Var}\{\eta(\nu_{p,h} - \nu_p)\} = \eta - \eta^2 \nu_{p,h} - \nu_p \\ & \eta^2 \text{Var}(\nu_{p,h}) - \eta \text{Cov}(\nu_{p,h}, \eta - \eta^2 \nu_{p,h} - \nu_p) = \text{Var}\{\eta - \eta^2 \nu_{p,h} - \nu_p\}. \end{aligned}$$

By employing the eff. $\eta^{-1} \nu_p f(\nu_p) = O(1)$

$$\eta^2 \text{Var}(\nu_{p,h}) = p^{-2} \nu_p^2 \text{Var}\{n^{-1} \sum_{t=1}^n I(Y_t > \nu_p)\} = p^{-2} n^{-1} b \nu_p^2 f(\nu_p) c_K = o(n^{-1} b) \quad \text{A.4}$$

we have $\eta \leq \eta_p$ then by using the fact that $\alpha_p \leq \alpha$ in consequence

$$E(\nu_{p,h} - \nu_p)^4 \leq C n^{-2} \quad \text{and} \quad E(\eta - \eta^2) = O(n^{-2-3})$$

Applying the Cauchy-Schwarz inequality and Lemma

$$\text{Var}\{\eta - \eta^2 \nu_{p,h} - \nu_p\} = O(n^{-2-3/2}) = o(n^{-1}) \quad \text{and} \quad \text{A.4}$$

$$\text{Cov}\{\eta(\nu_{p,h} - \nu_p), p^{-1}(\eta - \eta^2 \nu_{p,h} - \nu_p)\} = o(n^{-1}) \quad \text{A.4}$$

Combining A.4, A.4 and A.4

$$\text{Var}\{\eta(\nu_{p,h} - \nu_p)\} = p^{-2} \nu_p^2 \text{Var}\{\phi_2(\nu_p)\} = p^{-2} n^{-1} \nu_p^2 f(\nu_p) c_K = o(n^{-1}) \quad \text{A.4}$$

we have the following consequence on the left hand side of A.4 is

$$\begin{aligned} & \text{Cov}\left\{np^{-1} \sum_{t=1}^n Y_t G_h(\nu_p - Y_t), \eta(\nu_{p,h} - \nu_p)\right\} \\ & \text{Cov}\left\{np^{-1} \sum_{t=1}^n Y_t G_h(\nu_p - Y_t), \eta f^{-1}(\nu_p) n^{-1} \sum_{t=1}^n G_h(\nu_p - Y_t)\right\} = o(n^{-1}) \\ & np^2 p^{-1} \nu_p \left[\text{Cov}\{Y_t G_h(\nu_p - Y_t), G_h(\nu_p - Y_t)\} \right. \\ & \left. \sum_{k=1}^{n-1} -k/n \text{Cov}\{Y_1 G_h(\nu_p - Y_1), G_h(\nu_p - Y_{k+1})\} \right] = o(n^{-1}) \end{aligned}$$

hence $\text{Cov}\{Y_t G_h(\nu_p - Y_t), G_h(\nu_p - Y_t)\} = p - p^2 \nu_p - \nu_p f(\nu_p) c_K = o(1)$

$$\begin{aligned} & \text{Cov}\left\{np^{-1} \sum_{t=1}^n Y_t G_h(\nu_p - Y_t), np^{-1} \sum_{i=1}^n Y_i K_h(\nu_p - Y_i) \nu_{p,h} - \nu_p\right\} = \text{A.4} \\ & n^{-1} p^{-2} \nu_p \text{Cov}\{\phi_1(\nu_p), \phi_2(\nu_p)\} = n^{-1} p^{-2} \nu_p^2 f(\nu_p) c_K = o(1) \end{aligned}$$

Let A_1, A_2 and A_3 be non-negative random variables of $O(n^{-1})$ variance of A_1 and A_2 is

$$\text{Var}(\mu_p) = p^{-1}n^{-1}\sigma_0^2(p, n) = O(n^{-1}) \quad A_1$$

Let A_1, A_2 and A_3 be

Let A_1, A_2 and A_3 be $\nu_{p,h}$ can be used for A_1 and A_2 in the case of n and A_3 is the proof of the case

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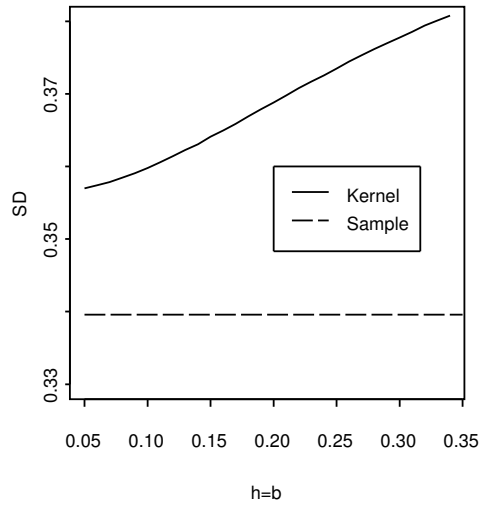
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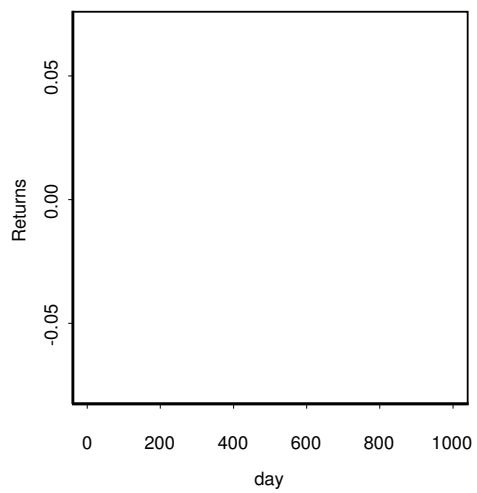
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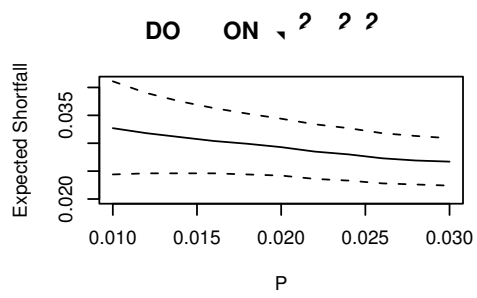
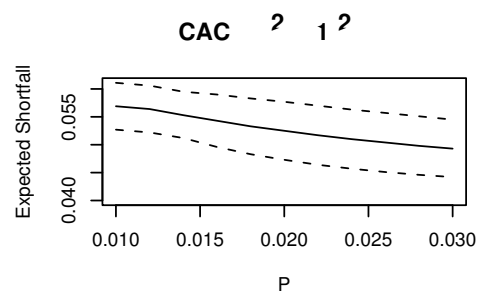
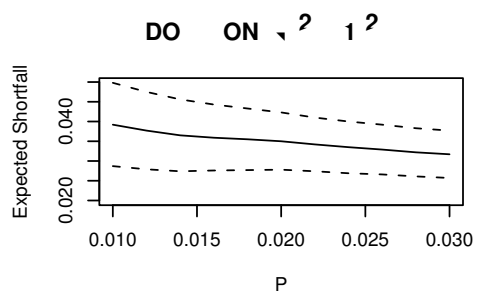
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(a) ES Estimates







T_{CAC}^* Estimates for $\nu_{0.01}$ $\mu_{0.01}$ and Standard Errors S.E

Ye	CAC		Do one	
	$\nu_{0.01}$	μ_p	$\nu_{0.01}$	μ_p
•				
•				
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